



## Auditor reputation and earnings management: International evidence from the banking industry

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### ABSTRACT

We examine the relation between auditor reputation and earnings management in banks using a sample of banks from 29 countries. In particular, we examine the implications of two aspects of auditor reputation, auditor type and auditor industry specialization, for earnings management in banks. We find that both auditor type and auditor industry specialization moderate benchmark-beating (loss-avoidance and just-meeting-or-beating prior year's earnings) behavior in banks. In addition, we find that once auditor type and auditor industry specialization are included in the same tests, only auditor industry specialization has a significant impact on constraining benchmark-beating behavior. In separate tests related to income-increasing abnormal loan loss provisions, we find that both auditor type and auditor expertise constrain income-increasing earnings management. Again, in joint tests, only auditor industry expertise has a significant impact on constraining income-increasing earnings management.

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### 1. Introduction

We examine the effect of auditor reputation on bank earnings management using an international sample of banks. Banks operate in a highly regulated environment in that they are monitored by Central Banks and other regulatory agencies (such as deposit insurance corporations). Consequently, auditor reputation may not be as important in constraining income-increasing earnings management in banks. On the contrary, if our finding establishes a negative association between auditor reputation and income-increasing earnings management, then auditor reputation likely is even more relevant for firms in other industries that are not subject to such direct regulatory scrutiny. To our knowledge no other study has examined how auditor reputation is related to earnings management in the international or the US banking industry.

Our main prediction is that auditor reputation (auditor type and auditor industry specialization) is negatively related to earnings management in banks even after controlling for several previously identified international institutional factors and bank monitoring

factors. Evidence of a negative relation between auditor reputation and earnings management may not be surprising for US banks because, in a high-litigation environment such as the US, high-reputation auditors have an incentive to maintain a high level of earnings quality to protect their reputation and legal exposure (Francis and Wang, 2008). Whether such a relation exists across different legal and institutional environments clearly is of interest. We are able to address this issue by analyzing an international sample of banks.

Prior research in banking has examined the relation between international institutional factors and bank monitoring variables and earnings management (Shen and Chih, 2005; Fonseca and González, 2008). Shen and Chih (2005) using earnings benchmark tests document that most banks manage their earnings. They also show that stronger investor protection and greater transparency in accounting disclosure reduce a bank's incentives to manage earnings. Fonseca and González (2008) focus on factors influencing income smoothing through loan loss provisions, the major bank accrual. They find that income smoothing is lower in jurisdictions with greater bank regulation and supervision. Interestingly, neither of these papers addresses the impact of auditing, an important external monitoring mechanism, on earnings management or income smoothing.

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In the auditing literature, [Becker et al. \(1998\)](#) report lower earnings management in industrial firms for clients of Big 5 auditors. [Krishnan \(2003\)](#) finds that firms audited by industry specialists report lower discretionary accruals, a commonly used proxy for earnings management in industrial firms. And, [Francis and Wang \(2008\)](#), using an international sample, report that earnings quality is higher for firms that use Big 4 auditors, but their result holds only for regimes with strong investor protection. A notable deficiency of these studies is that they exclude firms in banking and financial services.

In addition, auditing banks is more complex than auditing industrial firms. In its May 2006 report on large firm Public Company Accounting Oversight Board (PCAOB) inspection deficiency analysis, the American Institute of Certified Public Accountants' (AICPA, 2006) Center for Public Company Audit Firms finds that banks' loan loss allowance ranks number one among the various deficiencies found by inspectors. This indicates that auditing the loan loss allowance and the related loan loss provision are challenging tasks for auditors in general. High-reputation auditors have incentives to provide high-quality audits to avoid jeopardizing their reputation capital. Thus, auditor reputation is potentially important in assessing the adequacy of loan losses and mitigating earnings management incentives of bank managers.

In this study, we extend the research on benefits of auditor reputation to the banking industry. Specifically, we examine two aspects of auditor reputation. First, we investigate the implications of auditor type (Big 5 vs. non-Big 5 auditors) for constraining income-increasing earnings management in banks. A large body of empirical research documents that higher audit quality is associated with Big 5 auditors.<sup>1</sup> Relative to non-Big 5 auditors, Big 5 auditors have greater expertise, resources, and more importantly, market-based incentives (e.g., mitigating the risk of litigation and protecting their reputation capital) to constrain the tendency of their audit clients to engage in aggressive reporting. Consequently, we predict that earnings management will be lower for banks audited by Big 5 auditors.

Second, we examine the effect of auditor industry specialization on reducing earnings management in banks. Auditors who are specialists in the banking industry can better assess the adequacy of the loan loss provisions than non-specialist auditors. Prior research documents that auditor industry specialization enhances financial reporting quality and mitigates fraudulent financial reporting ([Johnson et al., 1991](#); [Carcello and Nagy, 2004](#); [Krishnan, 2003, 2005](#)). We measure auditor industry specialization/expertise by an auditor's industry market share.<sup>2</sup>

We employ three traditional proxies of earnings management, managing earnings to avoid losses, managing earnings to just-meet-or-beat the prior year's earnings, and an accrual-based proxy (abnormal loan loss provisions), to test the extent of income-increasing earnings management through bank loan loss provisions. By using three different tests (accruals- and non-accruals-based tests), we strengthen the validity and robustness of our results. Our loss-avoidance/just-meeting-or-beating prior year's earnings tests closely resemble the methodology used by [Beatty et al. \(2002\)](#) and [Altamuro and Beatty \(2010\)](#). Our proxy for abnormal loan loss provisions is based on prior banking research on loan loss provisions ([Wahlen, 1994](#); [Kanagaretnam et al., 2004](#)).

We use an international bank sample from the *BankScope* database representing 29 countries over the period 1993 to 2006 to test our hypotheses. We find in separate tests that both auditor type

and auditor expertise moderate benchmark-beating (loss-avoidance and just-meeting-or-beating prior year's earnings) behavior in banks. However, we find that once auditor type and industry expertise are included in the same tests, only auditor industry expertise has a significant impact on constraining benchmark-beating behavior. In separate tests related to income-increasing abnormal loan loss provisions, we find that both auditor type and auditor expertise constrain income-increasing earnings management. Again, in joint tests, only auditor industry expertise has a significant impact on constraining income-increasing earnings management. Overall we find that audit specialists constrain income-increasing earnings management in banks. Our results are robust to several sensitivity tests including alternate classification of audit specialists, controlling for self-selection, and different bin-widths for benchmark tests.

The rest of this paper is organized as follows. The next section develops the hypotheses. Section 3 explains the empirical models used for tests of earnings management. Section 4 describes the sample selection process. Section 5 discusses the results and Section 6 concludes the study.

## 2. Hypotheses

Using economic theory, [DeAngelo \(1981\)](#) argues that auditor size is a proxy for auditor reputation and audit quality. She reasons that brand-name auditors (i.e., Big 5 auditors) are better able to detect material misstatements in financial statements and more willing to report what they find than are other auditors (i.e., non-Big 5 auditors). Higher expertise is associated with Big 5 auditors because they not only have more resources but also devote more resources to specialized staff training, peer reviews, and investment in information technology than non-Big 5 auditors ([Craswell et al., 1995](#)). Similarly, higher independence is associated with Big 5 auditors because they have higher reputation capital at stake relative to non-Big 5 auditors. Loss of reputation, as Arthur Andersen learned the hard way, could put a Big 5 auditor out of business ([Huang and Li, 2009](#)). Litigation risk also motivates Big 5 auditors to remain independent. In the wake of the Enron-Andersen scandal PricewaterhouseCoopers, Deloitte and Touche, and Ernst and Young have resigned from more than 1200 clients to mitigate the risk of litigation ([Hindo, 2003](#)). In short, a higher audit quality is associated with the Big 5 auditors.

There is a large body of empirical research that documents that higher audit quality is associated with Big 5 auditors for industrial firms. [Teoh and Wong \(1993\)](#) observe higher earnings response coefficients for clients of Big 5 auditors relative to clients of non-Big 5 auditors, consistent with investors perceiving earnings to be of higher quality when the auditor is a brand-name auditor. [Becker et al. \(1998\)](#) report lower earnings management in industrial firms for clients of Big 5 auditors. Additionally, [Basu et al. \(2000\)](#) find higher levels of financial reporting conservatism (i.e., more timely recognition of bad news) for clients of Big 5 auditors. Although empirical evidence on auditor reputation and audit quality in the banking industry is limited, the economic incentives faced by the Big 5 auditors of banks are similar to those of other industries, i.e., preserving reputation capital and mitigating the risk of litigation.<sup>3</sup> In addition, auditor type may be of higher importance for industries such as banking, where information uncertainty is higher relative to industrial firms due to the greater complexity of banking operations and difficulty of assessing risk on the large portfolio of loans ([Autore et al., 2009](#)).

<sup>1</sup> We refer to the high-reputation, brand-name auditors as Big 5 (in fact Big 6, during the period 1993–1997, Big 5 after the merger between Coopers and Lybrand and Price Waterhouse in 1998, and currently Big 4 after the demise of Arthur Andersen in 2002) auditors throughout the paper for simplicity.

<sup>2</sup> In our discussions, we use auditor industry specialization and industry expertise interchangeably.

<sup>3</sup> [Kanagaretnam et al. \(2009\)](#) is an exception. They find a significant, positive association between the discretionary component of LLP and stock returns for banks audited by Big 5 auditors.

The above arguments suggest that earnings management in banks will be lower when the auditor is a Big 5 auditor. However, because banks operate in a highly regulated environment in that they are monitored by the Central Banks and other regulatory agencies (such as deposit insurance corporations), auditor reputation may not be as important in constraining income-increasing earnings management relative to industrial firms. Given this, we present our hypothesis in null form:

**Hypothesis 1.** Earnings management in banks is unrelated to whether the auditor is a Big 5 auditor.

Our second hypothesis relates to the linkage between auditor specialization in the banking industry and its impact on limiting earnings management behavior. While classifying auditors as Big 5 and non-Big 5 separates them in terms of reputation, it ignores an auditor's expertise in a given industry which may be an even more critical dimension of auditor reputation. Although a Big 5 auditor may serve clients in multiple industries, the auditor may not have a competitive advantage in all industries because developing a competitive advantage is very costly or not feasible due to a variety of reasons, including first-mover advantage enjoyed by other auditors, lack of economies of scale, limited human capital with industry expertise, and constrained economic resources. As a result, each of the Big 5 auditors or national auditors tends to dominate a select few industries from among the industries in their portfolio of clients.

Several studies have examined the benefits of auditor industry specialization or expertise on audit effectiveness. For example, *Bedard and Biggs (1991)* document that an auditor with experience in the manufacturing industries is better able to detect errors in a manufacturing client's data than an auditor without manufacturing experience. Similarly, *Wright and Wright (1997)* find that significant experience in the retailing industry contributes to increased detection of errors of clients in the retail industry. Other benefits of auditor industry specialization identified in prior research include mitigation of financial fraud (*Johnson et al., 1991; Carcello and Nagy, 2004*), reporting of lower discretionary accruals, a commonly used proxy for earnings management in industrial firms (*Krishnan, 2003*), and greater asymmetric timeliness of earnings which is a fundamental characteristic of conservative financial reporting (*Krishnan, 2005*).

While the above evidence indicates that the ability to detect material misstatements in financial statements is associated with auditor industry specialization, there also is evidence that specialist auditors attempt to protect their reputation capital through increased compliance with generally accepted auditing standards relative to non-specialist auditors (*O'Keefe et al., 1994*). In the banking industry, *Kanagaretnam et al. (2009)* find that once auditor type and industry expertise are separated, only auditor industry expertise has a significant impact on valuation of discretionary LLP. In summary, the collective evidence indicates that there are benefits to auditor industry specialization in terms of enhanced audit effectiveness and credibility of financial statements. However, as discussed earlier, the benefits of audit specialists may not be as pronounced in banking as in industrial firms because the banking industry is highly regulated. Accordingly, we present the following hypothesis stated in null form:

**Hypothesis 2.** Earnings management in banks is unrelated to whether the auditor is a specialist in the banking industry.

### 3. Measures of earnings management

Our general hypothesis is that auditor reputation (auditor type and auditor specialization) constrains bank earnings managements. We employ three traditional proxies of earnings management; managing earnings for loss-avoidance, managing earnings for just-meeting-or-beating prior year's earnings, and an accruals-based proxy, to

strengthen the validity/robustness of the results of our test of the extent of income-increasing earnings management.

#### 3.1. Managing earnings for loss-avoidance or to just-meet-or-beat prior year's earnings

*Beatty et al. (2002)* and *Altamuro and Beatty (2010)* report that bank managers have incentives to manage earnings for benchmark-beating behavior. We examine how auditor reputation (auditor type and auditor specialization) constrains this incentive. We focus on two earnings benchmarks: loss-avoidance (*LOSS\_AVOID*) and just-meeting-or-beating prior year's earnings (*JMBE*).<sup>4</sup> We include all available additional control variables (size, growth, loans, leverage, change in cash flow, and loan loss allowance) to be consistent with the above literature. We also include additional controls for country-specific variables (law enforcement index and financial system in the country), and bank-specific monitoring variables (regulatory restrictions on entry into banking, official supervisory power index, and private monitoring index), and estimate the following logit model:

$$\begin{aligned} \text{BENCHMARK}_t = & \alpha_0 + \alpha_1 \text{BIG5} + \alpha_2 \text{SPEC} + \alpha_3 \text{SIZE}_t \\ & + \alpha_4 \text{GROWTH}_t + \alpha_5 \text{LOANS}_t + \alpha_6 \text{LEV}_t \\ & + \alpha_7 \Delta \text{CASH\_FLOW}_t + \alpha_8 \text{ALLOW}_t \\ & + \alpha_9 \text{LAW\_ENF} + \alpha_{10} \text{BANK} + \alpha_{11} \text{BANKREG} \\ & + \alpha_{12} \text{OFFICIAL} + \alpha_{13} \text{MONITOR} + \\ & < \text{Year Controls} > + e, \end{aligned} \quad (1)$$

where

BENCHMARK	earnings benchmark indicators ( <i>LOSS_AVOID</i> or <i>JMBE</i> ), defined as follows: <i>LOSS_AVOID</i> is an indicator variable taking the value one if the bank has a small ROA (income before taxes scaled by total assets) in the interval between 0 and 0.002, and <i>JMBE</i> is an indicator variable taking the value one if the bank has a change in ROA (income before taxes scaled by total assets) from year $t - 1$ to year $t$ in the interval between 0 and 0.0005, zero otherwise
BIG5	an indicator variable that equals one if the auditor is a Big 5 auditor, zero otherwise
SPEC	an indicator variable that equals one if the auditor is a market leader in the banking industry for that particular country, zero otherwise
SIZE <sub>t</sub>	log of total assets at the end of the year
GROWTH <sub>t</sub>	the growth in total assets from the beginning to the end of year $t$
LOANS <sub>t</sub>	total loans scaled by total assets at the beginning of year $t$
LEV <sub>t</sub>	total equity divided by total assets at beginning of year $t$
ΔCASH_FLOW <sub>t</sub>	change in cash flows (income before taxes and loan loss provisions) from the beginning to the end of year $t$ scaled by total assets at the beginning of year $t$
ALLOW <sub>t</sub>	allowance for loan losses at the end of year $t$ , scaled by total assets at beginning of year $t$

<sup>4</sup> A recent survey of managers by *Graham et al. (2005)* finds that just-meeting-or-beating prior period's earnings is one of the most important benchmarks for corporate managers. In addition, *Burgstahler and Dichev (1997)* and *Degeorge et al. (1999)* provide empirical evidence indicating that loss-avoidance is also an important benchmark for managers.

Table (continued)

LAW_ENF	investor protection, proxied by the law enforcement index. This index is the mean score of three legal enforcement variables reported in LaPorta et al. (1998), and used in Leuz et al. (2003)
BANK	a bank system dummy, which equals 1 for countries whose financial system is bank-dominated and 0 for countries whose financial system is market-oriented, as per the classification of Demircug-Kunt and Levine (1999)
BANKREG	a measure of regulatory restrictions on entry into banking from Barth et al. (2001)
OFFICIAL	the official supervisory power index from Barth et al. (2001)
MONITOR	the private monitoring index from Barth et al. (2001)

The coefficients of interest are the coefficients of the two auditor reputation proxies: *BIG5* and *SPEC*. *BIG5* is an indicator variable that equals one if the auditor is a Big 5 auditor, zero otherwise. Auditor industry specialization/expertise is typically measured by an auditor's industry market share (e.g., Balsam et al., 2003; Krishnan, 2003) in each country. We calculate an auditor's market share as the fraction of a bank's total assets to the total assets of all banks in the same country. An auditor is considered a specialist if it has the largest market share in the banking industry for that particular country.<sup>5</sup> If higher auditor reputation reduces earnings management for loss-avoidance/just-meeting-or-beating prior year's earnings, the coefficients of *BIG5* and *SPEC* will be negative. On the other hand, if auditor reputation is not important in a highly regulated industry such as banking, the coefficients of *BIG5* and *SPEC* will not significantly differ from zero.

We include the two country-specific variables, *LAW\_ENF* and *BANK*, as control variables in all our regressions but do not offer directional predictions on the coefficients of these variables. We use the Law Enforcement Index (*LAW\_ENF*) to proxy for investor protection.<sup>6</sup> Leuz et al. (2003) find that *LAW\_ENF* is negatively associated with earnings management. Prior studies (e.g., Kwok and Tadesse, 2006; Purda, 2008) indicate that a firm's perceived risk (and the associated earnings quality due to earnings management) is influenced by whether the financial system in the country is bank- or market-based (*BANK*).

As suggested by Fonseca and González (2008), bank regulation and bank supervision may also affect the extent of earnings management. We use the measure of regulatory restrictions on entry to banking (*BANKREG*), the official supervisory power index (*OFFICIAL*), and the private monitoring index (*MONITOR*) developed by Barth et al. (2001) to proxy for bank regulation.<sup>7</sup> Since these vari-

<sup>5</sup> For each country, we identify the audit specialist for three distinct periods spanning our sample years. These three periods are: (1) 1993–1997, before the merger of Coopers and Lybrand and Price Waterhouse in 1998, (2) 1998–2001, after the merger of Coopers and Lybrand and Price Waterhouse, and (3) 2002–2006, after the demise of Arthur Andersen in 2002. This approach is consistent with Neal and Riley (2004) and Kwon et al. (2007).

<sup>6</sup> This index is the mean score of three legal enforcement variables reported in La Porta et al. (1998). The three variables are (1) the mean for 1980–1983 of a variable provided by Business International Corp., capturing the efficiency and integrity of the judicial system; (2) the mean for 1982–1995 of a rule of law variable obtained from International Country Risk; and (3) the mean for 1982–1995 of a corruption variable that assesses the corruption in government, obtained from International Country Risk.

<sup>7</sup> *BANKREG* ranges from 0 to 8, with higher values indicating more restrictions on entry into banking. *OFFICIAL* ranges from 0 to 14; it captures the power of supervisors to take prompt corrective action, to restructure and reorganize troubled banks, and to declare a troubled bank insolvent. *MONITOR* ranges from 0 to 7; it measures the extent of monitoring by outsiders such as international ratings agencies. Higher values of *OFFICIAL* and *MONITOR* indicate greater power of supervisors and greater private oversight.

ables are included as controls, we do not offer directional predictions on their coefficients.

### 3.2. Income-increasing abnormal loan loss provisions

We use a two-stage approach to examine the relation between auditor reputation and income-increasing earnings management through LLP. We first estimate the normal or nondiscretionary component of LLP by regressing LLP on beginning loan loss allowance, net loan charge-offs, change in total loans outstanding, total loans outstanding, nonperforming loans, loan categories, and controls for period and country effects using the following model<sup>8</sup>:

$$LLP = \lambda_0 + \lambda_1 BEGLLA + \lambda_2 LCO + \lambda_3 CHLOANS + \lambda_4 LOANS + \lambda_5 NPL + \lambda_6 DNPL + (\text{LOAN CATEGORIES}) + (\text{YEAR CONTROLS}) + (\text{COUNTRY CONTROLS}) + e \quad (2)$$

where

LLP	provisions for loan losses deflated by beginning total assets
BEGLLA	beginning loan loss allowance deflated by beginning total assets
LCO	net loan charge-offs deflated by beginning total assets
CHLOANS	change in total loans outstanding deflated by beginning total assets
LOANS	total loans outstanding deflated by beginning total assets
NPL	nonperforming loans deflated by beginning total assets
DNPL	an indicator variable that equals one if the value for NPL is missing, zero otherwise <sup>a</sup>
LOAN CATEGORIES	loans to municipalities/government ( <i>MUN</i> ), mortgages ( <i>MORT</i> ), hire-purchase/lease ( <i>LEASE</i> ), other loans ( <i>OTH</i> ), loans to group companies/associates ( <i>GRP</i> ), loans to other corporate ( <i>OCORP</i> ) and loans to banks ( <i>BK</i> ) all deflated by beginning total assets

<sup>a</sup> Since a large number of NPL observations are missing, we use the 'modified zero-order regression' method suggested by Maddala (1977) and Greene (2003) for estimating Eq. (2). This method substitutes a zero for missing values and adds an indicator variable coded one if the corresponding variable is missing.

The residuals from Eq. (2) are the abnormal or discretionary component of LLP, referred to as *ALLP*.

In the second stage, we test the association between proxies for auditor reputation and the absolute value of negative (income-increasing) *ALLP*. Negative *ALLP* are of particular interest because of their positive impact on reported earnings. We control for the following factors that prior research has documented to be associated with abnormal accruals (Ashbaugh et al., 2003): firm size, asset growth, level of past accruals, and performance. We use log of assets to measure size and prior period's LLP to proxy for level of past accruals. We represent performance by earnings before LLP. Our model is as follows:

$$ALLP = \delta_0 + \delta_1 BIG5 + \delta_2 SPEC + \delta_3 SIZE + \delta_4 GROWTH + \delta_5 PASTLLP + \delta_6 EBTP + \delta_7 LAW\_ENF + \delta_8 BANK + \delta_9 BANKREG + \delta_{10} OFFICIAL + \delta_{11} MONITOR + (\text{YEAR CONTROLS}) + \varepsilon, \quad (3)$$

<sup>8</sup> These variables have also been used in several prior studies (e.g., Wahlen, 1994; Kanagaretnam et al., 2004; Adams et al., 2009) to estimate the normal component of LLP.

where

ALLP	absolute value of negative abnormal loan loss provisions
PASTLLP	prior year's LLP divided by total assets at the beginning of the year
EBTP	earnings before taxes and loan loss provisions divided by total assets at the beginning of the year and all other variables are as previously defined

The coefficients of interest in Eq. (3) are the coefficients of *BIG5* and *SPEC*. A negative coefficient for both is consistent with our prediction of auditor reputation constraining income increasing earnings management. As previously discussed, if auditor reputation is not important in a highly regulated industry such as banking, the coefficients on *BIG5* and *SPEC* will not significantly differ from zero.

#### 4. Data description

We obtain financial data for the international (non-US) banks for the period 1993–2006 from the *BankScope* database. We select sample countries from the 48 countries listed in LaPorta et al. (1998). We drop nine countries (Ireland, Kenya, Pakistan, Sri Lanka, Zimbabwe, Ecuador, Indonesia, Jordan, and Venezuela) for which LaPorta et al. (1998) do not report legal enforcement variables needed to compute the law enforcement index (*LAW\_ENF*). We delete another 10 countries (Argentina, Belgium, Denmark, Egypt, Finland, Malaysia, Nigeria, Sweden, Switzerland, and Uruguay) due to missing bank-specific information (such as loan charge offs, loan loss provisions, loan loss allowance, etc). We retain the remaining 29 countries in our study. These include Australia, Austria, Brazil, Canada, Chile, Columbia, France, Germany, Greece, Hong Kong, India, Israel, Italy, Japan, Korea, Mexico, Netherlands, New Zealand, Norway, Peru, Philippines, Portugal, Singapore, South Africa, Spain, Taiwan, Thailand, Turkey, and the United Kingdom.

Our final sample consists of 6072 and 4232 bank-years for the earnings benchmark and abnormal loan loss provisions tests, respectively.<sup>9</sup> The sample for the benchmark tests is larger than the sample for the abnormal loan loss provisions test because of the less stringent data requirements. There is significant variation in the number of bank-year observations across countries due to differences in capital market development, country size, and the availability of complete financial accounting data. Except for the United Kingdom, most countries constitute less than ten percent of the total bank-years individually. In a sensitivity analysis (not reported in this paper), we find that our results are robust after excluding banks from the United Kingdom.

#### 5. Empirical results

The residuals from the regression models may be serially and/or cross-sectionally correlated. We therefore use OLS/logistic regressions with clustered robust errors to account for both serial and cross-sectional correlations (Rogers, 1993; Williams, 2000; Petersen, 2009). For all tests, we report Wald or *t*-statistics based on clustered standard errors after correcting for both serial and cross-sectional correlations in the residuals.<sup>10</sup>

<sup>9</sup> We delete each of the continuous control variables used in Eqs. (1)–(3) at the top and bottom one percent to remove extreme values.

<sup>10</sup> Petersen (2009) suggests that, in the presence of cross-sectional and time-series dependence, one dependence effect can be addressed parametrically (e.g., including time dummies for cross-sectional dependence) and then standard errors clustered on the other dependence effect (e.g., clustering by firms for time-series dependence) can be estimated. As we have more firm than year observations, we use year dummies and cluster by firms because having a large number of clusters results in standard errors that are less biased.

##### 5.1. Loss-avoidance and just-meeting-or-beating prior year's earnings tests

We report the results of the earnings benchmark tests in Table 1. Panel A presents the descriptive statistics of variables used in Eq. (1). On average, 6% of our sample banks report a small profit (i.e., they just avoided reporting a loss) and 7% of our sample banks report a small increase in earnings over the prior year (i.e., they just-meet-or-beat prior year's earnings). 74% (26%) of the banks are audited by Big 5 auditors (audit specialists). Panel B shows the correlations among the bank-specific variables used in the regression. Consistent with the argument that auditor reputation reduces earnings management, there is a negative correlation between banks audited by Big 5 and specialist auditors and loss-avoidance. However, there is a positive correlation between banks audited by Big 5 and specialist auditors and just-meeting-or-beating prior year's earnings. While these univariate results are interesting, they do not control for all other factors likely to affect the extent of earnings management. Hence, we rely on the multivariate analysis for making inferences.

We report the estimation results of the logistic regressions for the loss-avoidance and just-meeting-or-beating prior year's earnings tests in Table 2. A negative sign for  $\alpha_1$  and  $\alpha_2$ , the coefficients of *BIG5* and *SPEC*, will indicate that banks are less likely to manage earnings to avoid losses or just-meet-or-beat prior year's earnings when audited by high-reputation auditors. The first three columns provide the results for the loss-avoidance tests. In model (1), auditor reputation is proxied by whether or not the auditor is a Big 5 auditor. In model (2), auditor reputation is proxied by whether the auditor is a specialist in the banking industry for a particular country. In model (3), we include both proxies for auditor reputation. In model (1), we find a negative and significant (at the 10% level) coefficient of *BIG5*, indicating rejection of hypothesis 1. And, the negative and significant (at the 1% level) coefficient estimate for *SPEC* in model (2) indicates rejection of hypothesis 2. These results support the argument that high-reputation auditors constrain banks from managing earnings to avoid reporting losses. When we include both measures of auditor reputation in model (3), only the coefficient of *SPEC* maintains its significance level at 1%, suggesting that auditor industry specialization is more critical than auditor type in limiting earnings management.<sup>11</sup>

The next three columns provide the results of the JMBE tests. In model (1), we find that the coefficient of *BIG5* is negative and significant at the 10% level. In model (2), the coefficient of *SPEC* is negative and significant at the 1% level. In model (3), where we include both dimensions of auditor reputation, the coefficient of *SPEC* is negative and significant at the 10% level whereas the coefficient of *BIG5* is not significant.<sup>12</sup> These results suggest that banks audited by specialists (but not Big 5) are less likely to manage earnings to meet or beat the prior year's earnings.

Taken together, our results for the earnings benchmark tests indicate high reputation auditors, particularly auditors who are

<sup>11</sup> *LAW\_ENF* has significant negative association with *LOSS\_AVIOD*. This is consistent with Leuz et al. (2003), who find *LAW\_ENF* is negatively related to earnings management in industrial firms. *BANK* has significant positive association with *LOSS\_AVIOD*. This is consistent with bank-based financial systems having higher earnings management relative to market-based financial systems. Fonseca and González (2008) suggest that bank regulation and bank supervision may reduce the extent of earnings management. Since losses may induce closer scrutiny by regulators, an alternate prediction is that the higher bank regulation and bank supervision may increase the incentives for loss-avoidance. Consistent with this alternate prediction, *BANKREG*, *OFFICIAL*, and *MONITOR* are positively related to *LOSS\_AVIOD*.

<sup>12</sup> Consistent with Leuz et al. (2003), *LAW\_ENF* has significant negative association with *JMBE*. As before, *BANK* has significant positive association with *JMBE*. Consistent with Fonseca and González (2008), *BANKREG* has strong negative association with *JMBE*. As before, *OFFICIAL* is positively related to *JMBE*.

**Table 1**  
Descriptive statistics and correlations for Earnings Benchmark Test.

	Mean	Median	Q1	Q3	Std_dev					
<i>Panel A: Descriptive statistics</i>										
LOSS_AVOID	0.059	0.000	0.000	0.000	0.235					
JMBE	0.067	0.000	0.000	0.000	0.250					
BIG5	0.739	1.000	0.000	1.000	0.439					
SPEC	0.260	0.000	1.000	1.000	0.439					
SIZE	8.149	8.179	6.579	9.583	2.134					
GROWTH	0.145	0.096	0.020	0.192	0.319					
LOANS	0.631	0.652	0.426	0.842	0.300					
LEV	0.109	0.075	0.053	0.118	0.112					
ΔCASH_FLOW	0.003	0.001	−0.001	0.005	0.028					
ALLOW	0.010	0.010	0.004	0.015	0.007					
	LOSS_AVOID	JMBE	BIG5	SPEC	SIZE	GROWTH	LOANS	LEV	ΔCASH_FLOW	ALLOW
<i>Panel B: Correlations</i>										
LOSS_AVOID	1.000	0.062*	−0.089*	−0.148*	0.072*	−0.085*	−0.052*	−0.150*	−0.085*	0.003
JMBE		1.000	0.009	0.002	0.015	−0.027*	0.059*	−0.099*	−0.080*	−0.083*
BIG5			1.000	0.575*	−0.038*	0.002	−0.002	0.092*	−0.011	−0.169*
SPEC				1.000	−0.217*	0.108*	0.023	0.187*	0.025	−0.137*
SIZE					1.000	−0.009	0.030*	−0.583*	−0.010	0.053*
GROWTH						1.000	0.315*	−0.046*	0.253*	0.029*
LOANS							1.000	−0.136*	0.062*	0.143*
LEV								1.000	0.060*	0.127*
ΔCASH_FLOW									1.000	0.095*
ALLOW										1.000

Definitions for the firm-specific variables are as follow:

*LOSS\_AVOID* is an indicator variable that equals one if the bank has a small ROA (income before taxes scaled by total assets) in the interval between 0 and 0.002, zero otherwise; *JMBE* is an indicator variable that equals one if the bank has a change in ROA (income before taxes scaled by total assets) from year  $t - 1$  to year  $t$  in the interval between 0 and 0.0005, zero otherwise; *BIG5* is an indicator variable that equals one if the auditor is a Big 5 auditor, zero otherwise; *SPEC* is an indicator variable that equals one if the auditor is a market leader in the industry, zero otherwise; *SIZE* is the log of total assets at the beginning of the year; *GROWTH* is the growth in total assets from the beginning to the end of the year; *LOANS* is total loans scaled by total assets at the beginning of the year; *LEV* is total equity divided by total assets at beginning of the year; *ΔCASH\_FLOW* is the change in cash flows (income before taxes and before loan loss provisions) from the beginning to the end of the year scaled by total assets at the beginning of the year; and *ALLOW* is the allowance for loan losses at the end of the year, scaled by total assets at the beginning of the year.

\* Significance at 5% level (two-tailed).

industry specialists, help to constrain earnings management undertaken by banks.

## 5.2. Income-increasing abnormal loan loss provisions test

We report the results for the abnormal loan loss provisions tests in Tables 3–5. We use a two-stage approach to investigate the effect of auditor reputation on abnormal loan loss provisions. The sample size for the first stage regression is 7680 bank-years with all data available to estimate abnormal LLP. The sample for our second stage is reduced to 4232 bank-years due to missing auditor information in the *BankScope* database. Table 3 reports the results of the first-stage regression for estimating abnormal LLP. Consistent with prior studies (e.g., Kanagaretnam et al., 2009), *BEGLLA* is negatively associated with *LLP* since a higher initial loan loss allowance will require a lower *LLP* in the current period. As expected, *LCO*, *LOANS* and *NPL* are positively associated with *LLP*, consistent with the evidence reported in Kanagaretnam et al. (2004). The residuals from Eq. (2) represent the abnormal component of LLP.

We are primarily interested in how auditor reputation may affect income-increasing earnings management and hence we report in Table 4 the regression results for the absolute value of the 2442 bank-years with negative (income-increasing) *ALLP* values. Of interest are the coefficients  $\delta_1$  and  $\delta_2$  of *BIG5* and *SPEC*. A negative sign for  $\delta_1$  and  $\delta_2$  suggests less income-increasing earnings management when auditor reputation is high. Recall that we are using absolute values, so that smaller values of *ALLP* indicate less income-increasing earnings management. Consistent with expectations, the coefficient  $\delta_1$  in model (1), where auditor reputation is measured by *BIG5*, is negative and significant at the 10% level. In model (2), where auditor reputation is measured by *SPEC*, the coef-

ficient  $\delta_2$  is negative and significant at the 1% level. In model (3) where we include both measures of auditor reputation in the model, the coefficient of *BIG5* is negative but no longer significant at conventional levels and the coefficient of *SPEC* is negative and significant at the 1% level.<sup>13</sup> These results support our contention that high-reputation auditors constrain income-increasing earnings management by banks.<sup>14</sup>

Based on our research design, the most likely area to find earnings management is in the area of loss-avoidance or just-meeting-or-beating the prior year's results. This is also when the reputation of the audit firm on the income-increasing abnormal loan loss provision (*ALLP*) would be expected to be the greatest.<sup>15</sup> To test this conjecture, we add *LOSS\_AVOID* and *JMBE* and their interactions with *BIG5* and *SPEC* to model (3). In Table 5, we report the regression results for this expanded model (3). Addition of these variables does not alter our main results on the effects of *BIG5* and *SPEC* in reducing the income-increasing *ALLP* reported in Table 4. More interestingly, the sum of the coefficients of *BIG5* and *BIG5 \* JMBE*, *SPEC* and *SPEC \* LOSS\_AVOID*, and *SPEC \* JMBE*, are all negative and significant at the 5% level or better. This is consistent with our conjecture that the influence of auditor reputation will be greatest in constraining income-increasing earnings management through *ALLP* when banks have incentives for benchmark beating.

<sup>13</sup> Consistent with prior research, *LAW\_ENF* is significantly negatively associated with absolute value of negative (income-increasing) *ALLP*. And, consistent with Fonseca and Gonzalez (2008), *OFFICIAL* and *MONITOR* are negatively related to income-increasing *ALLP*. However, contrary to expectations, *BANK* is strongly negatively associated with income-increasing *ALLP*.

<sup>14</sup> In untabulated results, we find no relation between our measures of auditor reputation and income-decreasing earnings management by banks.

<sup>15</sup> We thank an anonymous referee for bringing this important issue to our attention.

**Table 2**  
Regression results for the loss-avoidance and just-meet-or-beat prior year's Earnings Benchmark Tests.

Variable	Coeff.	Loss-avoidance test			Just-meeting-or-beating prior year's earnings		
		Model (1)	Model (2)	Model (3)	Model (1)	Model (2)	Model (3)
Intercept	$\alpha_0$	-3.929 (10.24)***	-3.268 (7.24)***	-3.289 (7.35)***	-1.598 (2.67)*	-1.376 (1.99)	-1.376 (2.00)
BIG5	$\alpha_1$	-0.289 (2.82) <sup>†</sup>		-0.075 (0.14)	-0.265 (2.85) <sup>†</sup>		-0.084 (0.19)
SPEC	$\alpha_2$		-0.661 (8.78)***	-0.619 (5.79)***		-0.483 (6.36)***	-0.436 (3.59) <sup>†</sup>
SIZE	$\alpha_3$	-0.019 (0.23)	-0.036 (0.82)	-0.034 (0.72)	-0.085 (5.62)**	-0.096 (7.05)***	-0.094 (6.71)***
GROWTH	$\alpha_4$	-0.305 (0.95)	-0.271 (0.74)	-0.269 (0.73)	-1.191 (20.21)***	-1.164 (18.97)***	-1.167 (18.91)***
LOANS	$\alpha_5$	-1.506 (18.69)***	-1.425 (17.00)***	-1.434 (17.29)**	0.992 (14.57)***	1.046 (16.26)***	1.042 (16.17)***
LEV	$\alpha_6$	-8.886 (14.23)***	-8.885 (14.24)***	-8.828 (14.22)***	-7.125 (28.30)***	-7.070 (28.22)***	-7.047 (28.26)***
$\Delta$ CASH_FLOW	$\alpha_7$	-1.896 (0.33)	-1.869 (0.28)	-1.877 (0.29)	-5.293 (14.83)***	-5.482 (15.06)***	-5.492 (15.10)***
ALLOW	$\alpha_8$	-0.505 (0.01)	-0.174 (0.01)	-0.452 (0.01)	-37.059 (12.09)***	-36.464 (11.70)***	-36.629 (11.68)***
LAW_ENF	$\alpha_9$	-0.022 (9.54)***	-0.021 (7.44)***	-0.022 (7.74)***	-0.016 (5.69)**	-0.015 (5.17)**	-0.015 (5.25)**
BANK	$\alpha_{10}$	0.938 (12.96)***	0.734 (6.69)***	0.742 (6.93)***	0.432 (6.53)***	0.375 (6.38)**	0.376 (4.42)**
BANKREG	$\alpha_{11}$	0.108 (11.65)***	0.075 (4.43)**	0.073 (4.28)**	-0.171 (38.92)***	-0.188 (38.46)***	-0.189 (40.19)***
OFFICIAL	$\alpha_{12}$	0.054 (1.99)	0.079 (3.10)*	0.078 (3.03)*	0.130 (15.62)***	0.145 (14.94)***	0.144 (14.77)***
MONITOR	$\alpha_{13}$	0.379 (4.61)**	0.338 (3.31)*	0.348 (3.60)*	-0.049 (0.22)	-0.054 (0.23)	-0.049 (0.18)
Year Controls		Yes	Yes	Yes	Yes	Yes	Yes
N		6072	6072	6072	6072	6072	6072
Likelihood ratio		313.87***	323.92***	324.15***	220.71***	225.69***	225.95***

The regression model is:

$$\text{BENCHMARK} = \alpha_0 + \alpha_1 \text{BIG5} + \alpha_2 \text{SPEC} + \alpha_3 \text{SIZE}_t + \alpha_4 \text{GROWTH} + \alpha_5 \text{LOANS}_t + \alpha_6 \text{LEV}_t + \alpha_7 \Delta \text{CASH\_FLOW}_t + \alpha_8 \text{ALLOW}_t + \alpha_9 \text{LAW\_ENF} + \alpha_{10} \text{BANK} + \alpha_{11} \text{BANKREG} + \alpha_{12} \text{OFFICIAL} + \alpha_{13} \text{MONITOR} + (\text{Year Controls}) + e,$$

where *BENCHMARK* is defined as *LOSS\_AVOID* or *JMBE*. *BANK* is a bank system dummy variable that equals one for countries whose financial system is bank-dominated and zero for countries whose financial system is market-oriented, as per the classification of Demiguc-Kunt and Levine (1999); *LAW\_ENF* is the law enforcement index. The index is the mean score of three legal enforcement variables reported in LaPorta et al. (1998), and used in Leuz et al. (2003). The three variables are (1) the mean for 1980–1983 of a variable provided by Business International Corp., capturing the efficiency and integrity of the judicial system; (2) the mean for 1982–1995 of a rule of law variable obtained from International Country Risk; and (3) the mean for 1982–1995 of a corruption variable that assesses the corruption in government, obtained from International Country Risk; *BANKREG* is a measure of regulatory restrictions on entry into banking; *OFFICIAL* is the official supervisory power index; *MONITOR* is the private monitoring index. All other variables are defined in the footnotes of Table 1. We run the logistic regression clustered by firm, and with year dummies. To conserve space, we do not report the coefficient estimates for the year dummies. For each variable, we report the regression coefficient, followed by the Wald statistic in parentheses.

<sup>†</sup> Significance at 10% level (two-tailed).

\*\* Significance at 5% level (two-tailed).

\*\*\* Significance at 1% level (two-tailed).

### 5.3. Sensitivity checks

In the main analysis, we define the industry market leader for each country as the specialist auditor. We use the Neal and Riley (2004) criteria, which allow more than one auditor to be a specialist in the banking industry, as an alternative definition. Specifically, we define an auditor with a large industry market share (based on bank assets) as a specialist (*SPEC1*) for each country. We consider an auditor to have a large market share if the auditor has at least a 20% share of the industry for the 1993–1997 period, a 24% share of the industry for the 1998–2001 period, and a 30% share of the industry for the 2002–2005 period.<sup>16</sup> For the loss-avoidance test, in model (2), the coefficient of *SPEC1* is negative and significant. In

model (3), the coefficient for *BIG5* is insignificant but the coefficient for *SPEC1* is negative and significant. For the just-meeting-or-beating prior year's earnings test, the coefficient for *SPEC1* is negative and significant in model (2). In model (3), the coefficient for *BIG5* is insignificant but the coefficient for *SPEC1* is negative and significant. Overall, our results are robust to this alternative definition of auditor specialization.

Second, we conduct a sub-sample analysis using only US banks. As discussed earlier, we omit US banks because they operate in a highly litigious environment that likely differs from the environment in other countries. Because US banks were generally very profitable during our sample period, we have very few observations in the bin-width for the loss-avoidance and just-meeting-or-beating prior year's earnings tests. In separate tests related to absolute value of income-increasing abnormal loan loss provisions, we find that both auditor type and auditor expertise are negatively related to income-increasing earnings management.

Third, we control for potential endogeneity of auditor choice and banks' earnings management. Banks with high earnings quality may systematically choose high-reputation auditors and

<sup>16</sup> Following Neal and Riley (2004), we employ a cutoff for "large" market shares of  $(1/N) * 1.2$ , where *N* is the number of big international audit firms. The largest firms are the Big 6, during the period 1993–1997, the Big 5 after the merger between Coopers and Lybrand and Price Waterhouse in 1998, and the Big 4 after the demise of Arthur Andersen in 2002. This measure includes all firms that cross the 20%/24%/30% thresholds and is denoted as *SPEC1*.

**Table 3**  
Stage-one regression in estimating abnormal loan loss provisions.

Variable	Coeff.	Estimate
Intercept	$\lambda_0$	-0.001 (-0.54)
BEGLLA	$\lambda_1$	-0.009 (-3.69)***
LCO	$\lambda_2$	0.022 (5.68)***
CHLOANS	$\lambda_3$	-0.006 (-6.64)***
LOANS	$\lambda_4$	0.010 (15.03)***
NPL	$\lambda_5$	0.077 (2.51)***
DNPL	$\lambda_6$	0.008 (0.98)
MUN	$\lambda_7$	-0.007 (-1.49)
MORT	$\lambda_8$	-0.006 (-8.65)***
LEASE	$\lambda_9$	-0.004 (-2.72)***
OTH	$\lambda_{10}$	-0.001 (-1.34)
GRP	$\lambda_{11}$	-0.019 (-6.04)***
OCORP	$\lambda_{12}$	-0.037 (-0.43)
BK	$\lambda_{13}$	-0.013 (-0.23)
Year dummies		Yes
Country dummies		Yes
N		7680
Adjusted R <sup>2</sup>		9.58

We report the results for the following stage one regression model:

$$LLP_{it} = \lambda_0 + \lambda_1 BEGLLA + \lambda_2 LCO + \lambda_3 CHLOANS + \lambda_4 LOANS + \lambda_5 NPL + \lambda_6 DNPL + \langle LOAN \text{ CATEGORIES} \rangle + \langle YEAR \text{ CONTROLS} \rangle + \langle COUNTRY \text{ CONTROLS} \rangle + e_{it}$$

where *LLP* is the provisions for loan losses; *BEGLLA* is the beginning loan loss allowance; *LCO* is net loan charge-offs; *CHLOANS* is the change in total loans outstanding; *LOANS* is total loans outstanding; *NPL* is the non-performing loans; and *DNPL* is a dummy variable that equals one if *NPL* is missing, zero otherwise. These variables are deflated by beginning total assets. *LOAN CATEGORIES* is an indicator variable for different type of loans. They are defined as: Loans to Municipalities/Government (*MUN*); Mortgages (*MORT*), HP/Lease (*LEASE*), Other Loans (*OTH*), Loans to Group companies/Associates (*GRP*), Loans to Other Corporate (*OCORP*), Loans to Banks (*BK*). For each variable, we report the regression coefficient, followed by the *t*-statistic in parentheses.

\*\*\* Significance at 1% level (two-tailed).

high-reputation auditors may likewise prefer to audit banks with high earning quality. We employ the Heckman (1979) two-stage procedure to address this concern.<sup>17</sup> We first estimate a probit model of auditor choice (Big 5 vs. non-Big 5) to derive the Inverse Mills Ratio (*IMR*). We are not aware of a model of auditor choice for banks. Therefore, we develop the following model that relates auditor choice to bank performance, size, and risk<sup>18</sup>:

$$BIG5 = \alpha_0 + \alpha_1 ROA + \alpha_2 ROA * LOSS + \alpha_3 LnLOAN + \alpha_4 CHLOAN + \alpha_5 TCAPITAL + \alpha_6 LOANRATIO + \alpha_7 NPLRATIO + \varepsilon, \quad (4)$$

<sup>17</sup> Francis and Lennox (2008) discuss the problems with operationalizing the Heckman (1979) two-stage procedure in an accounting context. Following their suggestion, we exclude several variables that appear in the first stage model (for example *LnLOAN* and *NPLRATIO*) from our second stage model. We also check for multicollinearity when including *IMR* in the second stage. The variance inflation factors do not indicate problems with multicollinearity.

<sup>18</sup> The sample size is reduced due to missing information for capital ratio and non-performing loans.

where

<i>BIG5</i>	an indicator variable that equals 1 for Big 5 clients and 0 for non- Big 5 clients
<i>ROA</i>	net income over lagged total assets
<i>LOSS</i>	an indicator variable that equals 1 the firm has a loss and 0 otherwise
<i>LnLOAN</i>	natural log of total loans outstanding
<i>CHLOAN</i>	change in total loans outstanding, scaled by beginning assets
<i>TCAPITAL</i>	risk adjusted total capital ratio (at the year-end)
<i>LOANRATIO</i>	total loans outstanding divided by total assets
<i>NPLRATIO</i>	nonperforming loans divided by total loans

*ROA* captures performance and is expected to be positively related to *BIG5*. By including *ROA \* LOSS* in the model, we allow the coefficient of *ROA* to differ across profit and loss firms. *LnLOAN* is a proxy for bank size. Since larger banks may prefer Big 5 auditors, we predict a positive coefficient for *LnLOAN*. We use several measures of risk, including *CHLOAN*, *TCAPITAL*, *NPLRATIO* and *LOANRATIO*. We predict positive coefficients for *CHLOAN*, *TCAPITAL*, *LOANRATIO*, and *NPLRATIO* as banks with greater actual or perceived risks may opt for a Big 5 auditor to enhance the credibility of their financial reports. The (untabulated) results for our first-stage regression indicate that *ROA*, *LnLOAN*, *CHLOAN*, *LOANRATIO*, *NPLRATIO*, and *TCAPITAL* are positively and significantly associated with *BIG5* while *ROA \* LOSS* is negatively and significantly associated with *BIG5*.

We then re-estimate Eqs. (1) and (3) with *IMR* as an additional independent variable. In the second stage the coefficient of *IMR* in all our models is significant at the 1% level suggesting that self-selection may be a problem in our analysis. Despite the inclusion of *IMR* in the regression model, the results of the second-stage regression are similar to those reported in Tables 2 and 4.

Our fourth sensitivity test is related to the benchmark-beating test. In the loss-avoidance test, we consider a bank's attempt to avoid losses if its earnings are between 0 and 0.002. As a robustness check, we use alternative benchmarks: *SMALL\_PROFIT* is coded one if *ROA* is [0, 0.001] or [0, 0.003]. Our results are similar to those reported in Table 2. In the *JMBE* test, we use the alternative bin-width criterion of [0, 0.0001] or [0, 0.0008] in scaled change in earnings instead of [0, 0.0005] to classify firms as just-meeting-or-beating their earnings benchmark. Again, we obtain similar results.

Our fifth sensitivity test is related to the banks audited by Arthur Andersen. We re-estimate the models after excluding former Andersen clients (194 bank-year observations for the benchmark tests, and 62 bank-year observations for the abnormal loan loss provision test). Overall, our results are robust to the exclusion of these former Andersen clients. We also find that there is no difference in earnings management between Andersen and non-Andersen Big 5 clients. This is consistent with Cahan and Zhang (2006) who find no significant differences between the US clients of Andersen and those of the other Big-5 auditors in terms of unadjusted and performance-adjusted abnormal accruals.

Finally, we run additional tests by replacing the country level institutional variables with fixed effects for each country. Except for the *JMBE* test, our results are similar to those reported in the paper.

## 6. Summary and conclusions

Given the importance of banking to national and global economies, there is surprisingly little evidence on the implications of auditing for banks' earnings quality. Banks are very different from industrial firms, and given the recent, heightened concern with the



**Table 4**

Association between absolute value of income-increasing ALLP and auditor reputation.

Variable	Coeff.	Model (1)	Model (2)	Model (3)
Intercept	$\delta_0$	0.984 (7.31)***	0.992 (7.40)***	0.973 (7.22)***
BIG5	$\delta_1$	-0.038 (-1.81)*		-0.026 (-1.24)
SPEC	$\delta_2$		-0.065 (-3.07)***	-0.060 (-2.80)***
SIZE	$\delta_3$	-0.022 (-4.52)***	-0.021 (-4.43)***	-0.021 (-4.35)***
GROWTH	$\delta_4$	0.054 (2.06)**	0.056 (2.16)**	0.055 (2.11)**
PASTLLP	$\delta_5$	-1.143 (-1.40)	-1.078 (-1.33)	-1.142 (-1.41)
EBTP	$\delta_6$	0.140 (0.39)	0.047 (0.13)	0.081 (0.23)
LAW_ENF	$\delta_7$	-0.001 (-2.06)**	-0.002 (-2.44)**	-0.002 (-2.26)**
BANK	$\delta_8$	-0.066 (-3.01)***	-0.066 (-3.06)***	-0.072 (-3.29)***
BANKREG	$\delta_9$	-0.001 (-0.08)	-0.005 (-0.53)	-0.003 (-0.26)
OFFICIAL	$\delta_{10}$	-0.011 (-3.16)***	-0.011 (-3.13)***	-0.011 (-3.16)***
MONITOR	$\delta_{11}$	-0.057 (-3.76)***	-0.056 (-3.66)***	-0.053 (-3.47)***
Year controls		Yes	Yes	Yes
N		2407	2407	2407
Adj. R <sup>2</sup>		7.45	7.93	8.01

The table reports the results for the following stage-two regression model:

$$\text{ALLP} = \delta_0 + \delta_1 \text{BIG5} + \delta_2 \text{SPEC} + \delta_3 \text{SIZE} + \delta_4 \text{GROWTH} + \delta_5 \text{PASTLLP} + \delta_6 \text{EBTP} \\ + \delta_7 \text{LAW\_ENF} + \delta_8 \text{BANK} + \delta_9 \text{BANKREG} + \delta_{10} \text{OFFICIAL} + \delta_{11} \text{MONITOR} \\ + (\text{YEAR CONTROLS}) + \varepsilon.$$

The dependent variable is the absolute values of the negative (income-increasing) abnormal loan loss provisions. The definitions for the variables are provided in footnotes of Tables 1 and 3. The definitions of the institutional variables are provided in the footnotes of Table 2. We estimate the regression clustered by firm, and with year dummies. For ease of presentation, we multiply the coefficient estimates by 100. To conserve space, we do not report the coefficient estimates for the year dummies. For each variable, we report the regression coefficient, followed by the *t*-statistic in parentheses.

\* Significance at 10% level, two-tailed.

\*\* Significance at 5% level, two-tailed.

\*\*\* Significance at 1% level, two-tailed.

quality of banks' reported earnings following the meltdown in this sector, a study of the effect of auditor reputation on earnings management in the banking industry is of considerable interest to regulators and investors.

We examine the relation between auditor reputation and earnings management in banks using a sample of banks from 29 countries. We hypothesize that high-reputation auditors will constrain income-increasing earnings management in banks. In particular, we examine the implications of two aspects of auditor reputation, auditor type (Big 5 vs. non-Big 5 auditors) and auditor industry specialization for earnings management in banks.

We find in separate tests that both auditor type and auditor expertise constrain earnings benchmark-beating (loss-avoidance and just-meeting-or-beating prior year's earnings) behavior in banks. However, we find that once auditor type and industry expertise are included in the same tests, only auditor industry expertise has a significant impact on constraining benchmark-beating behavior. In tests related to income-increasing abnormal loan loss provisions, we find in separate tests that both auditor type and auditor expertise constrain income-increasing earnings management. Again, in joint tests, only auditor industry expertise has a significant impact on constraining income-increasing earnings management.

**Table 5**

Association between Absolute value of income-increasing ALLP, loss-avoidance, and just-meet-or-beat prior year's earnings, and auditor reputation.

Variable	Coef.	Model (1)	Model (2)	Model (3)
Intercept	$\delta_0$	0.963 (8.06)***	0.983 (8.14)***	0.965 (8.07)***
BIG5	$\delta_1$	-0.039 (-1.79)*		-0.028 (-1.30)
SPEC	$\delta_2$		-0.057 (-2.57)***	-0.052 (-2.30)**
SIZE	$\delta_3$	-0.022 (-4.65)***	-0.022 (-4.55)***	-0.022 (-4.47)***
GROWTH	$\delta_4$	0.055 (2.10)**	0.056 (2.18)**	0.055 (2.12)**
PASTLLP	$\delta_5$	-1.267 (-1.57)	-1.192 (-1.50)	-1.272 (-1.59)
EBTP	$\delta_6$	0.119 (0.33)	0.032 (0.09)	0.066 (0.19)
LAW_ENF	$\delta_7$	-0.001 (-1.90)*	-0.002 (-2.32)**	-0.001 (-2.10)**
BANK	$\delta_8$	-0.067 (-3.04)***	-0.066 (-3.03)***	-0.072 (-3.28)***
BANKREG	$\delta_9$	0.001 (0.03)	-0.005 (-0.49)	-0.002 (-0.17)
OFFICIAL	$\delta_{10}$	-0.010 (-3.04)***	-0.010 (-3.08)***	-0.010 (-3.05)***
MONITOR	$\delta_{11}$	-0.056 (-3.73)***	-0.056 (-3.66)***	-0.053 (-3.49)***
LOSS_AVOID	$\delta_{12}$	-0.028 (-0.80)	0.020 (0.66)	-0.028 (-0.81)
JMBE	$\delta_{13}$	0.014 (0.42)	-0.045 (-2.13)**	0.013 (0.42)
BIG5 * LOSS_AVOID	$\delta_{14}$	0.075 (1.41)		0.093 (1.60)
BIG5 * JMBE	$\delta_{15}$	-0.097 (-2.55)***		-0.091 (-2.18)**
SPEC * LOSS_AVOID	$\delta_{16}$		-0.122 (-2.65)***	-0.161 (-2.75)***
SPEC * JMBE	$\delta_{17}$		-0.038 (-1.11)	-0.008 (-0.20)
Year Controls		Yes	Yes	Yes
N		2407	2407	2407
Adj. R <sup>2</sup>		7.76	8.01	8.26

The table reports the results for the following stage-two regression model:

$$\text{ALLP} = \delta_0 + \delta_1 \text{BIG5} + \delta_2 \text{SPEC} + \delta_3 \text{SIZE} + \delta_4 \text{GROWTH} + \delta_5 \text{PASTLLP} + \delta_6 \text{EBTP} \\ + \delta_7 \text{LAW\_ENF} + \delta_8 \text{BANK} + \delta_9 \text{BANKREG} + \delta_{10} \text{OFFICIAL} + \delta_{11} \text{MONITOR} \\ + \delta_{12} \text{LOSS\_AVOID} + \delta_{13} \text{JMBE} + \delta_{14} \text{BIG5} * \text{LOSS\_AVOID} + \delta_{15} \text{BIG5} * \text{JMBE} \\ + \delta_{16} \text{SPEC} * \text{LOSS\_AVOID} + \delta_{17} \text{SPEC} * \text{JMBE} + (\text{YEAR CONTROLS}) + \varepsilon.$$

The definitions for the variables are provided in footnotes of Tables 1–3. We estimate the regression clustered by firm, and with year dummies. For ease of presentation, we multiply the coefficient estimates by 100. To conserve space, we do not report the coefficient estimates for the year dummies. For each variable, we report the regression coefficient, followed by the *t*-statistic in parentheses.

\* Significance at 10% level, two-tailed.

\*\* Significance at 5% level, two-tailed.

\*\*\* Significance at 1% level, two-tailed.

Our results show that even in a highly regulated industry such as banking, auditor reputation has an important role in constraining income-increasing earnings management. Moreover, in an international banking context, our study can be regarded as documenting an important external monitoring mechanism in addition to previously identified international institutional factors and bank monitoring factors that constrains earnings management in banks.

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## References

- Adams, B., Carow, K.A., Perry, T., 2009. Earnings management and initial public offerings: The case of the depository industry. *Journal of Banking and Finance* 33, 2363–2372.
- AICPA, 2006. Large Firm PCAOB Inspection Deficiency Analysis. <[http://www.aicpa.org/cpcf/download/PCAOB\\_Inspection\\_Deficiency\\_Analysis\\_January2006.pdf](http://www.aicpa.org/cpcf/download/PCAOB_Inspection_Deficiency_Analysis_January2006.pdf)>.
- Altamuro, J., Beatty, A., 2010. How does internal control regulation affect financial reporting? *Journal of Accounting and Economics* 49 (1–2), 58–74.
- Ashbaugh, H., LaFond, R., Mayhew, B., 2003. Do nonaudit services compromise auditor independence? Further evidence. *The Accounting Review* 78 (3), 611–639.
- Autore, D.M., Billingsley, R.S., Schneller, M.I., 2009. Information uncertainty and auditor reputation. *Journal of Banking and Finance* 33 (12), 183–192.
- Balsam, S., Krishnan, J., Yang, J., 2003. Auditor industry specialization and earnings quality. *Auditing: A Journal of Practice & Theory* (September), 71–97.
- Barth, J.R., Caprio, G., Levine, R., 2001. The Regulation and Supervision of Banks Around the World: A New Database. Working Paper, World Bank.
- Basu, S., Hwang, L., Jan, C., 2000. Differences in Conservatism Between Big Eight and Non-Big Eight Auditors. Working Paper, City University of New York and California State University.
- Beatty, A.L., Ke, B., Petroni, K.R., 2002. Earnings management to avoid earnings declines across publicly and privately held banks. *The Accounting Review* 77 (3), 547–570.
- Becker, C., DeFond, M., Jiambalvo, J., Subramanyam, K., 1998. The effect of audit quality on earnings management. *Contemporary Accounting Research* 15 (Spring), 1–24.
- Bedard, J.C., Biggs, S.F., 1991. The effect of domain-specific experience on evaluation of management representations in analytical procedures. *Auditing: A Journal of Practice & Theory* (Supplement), 77–90.
- Burgstahler, D., Dichev, I., 1997. Earnings management to avoid earnings decreases and losses. *Journal of Accounting and Economics* 24 (1), 99–126.
- Cahan, S.F., Zhang, W., 2006. After Enron: Auditor conservatism and ex-Andersen clients. *The Accounting Review* 81 (1), 49–82.
- Carcello, J., Nagy, A., 2004. Client size, auditor specialization and fraudulent financial reporting. *Managerial Auditing Journal* 19 (5), 651–668.
- Craswell, A., Francis, J., Taylor, S., 1995. Auditor brand name reputations and industry specializations. *Journal of Accounting and Economics* 20 (December), 297–322.
- DeAngelo, L., 1981. Auditor size and auditor quality. *Journal of Accounting and Economics* 3 (December), 183–199.
- DeGeorge, F., Patel, J., Zeckhauser, R., 1999. Earnings management to exceed thresholds. *The Journal of Business* 72, 1–33.
- Demircug-Kunt, A., Levine, R., 1999. Bank-based and Market-based Financial Systems: Cross-country Comparisons. Working Paper, World Bank.
- Fonseca, A., González, F., 2008. Cross-country determinants of bank income smoothing by managing loan-loss provisions. *Journal of Banking and Finance* 32 (2), 217–218.
- Francis, J., Lennox, C., 2008. Selection Models in Accounting Research. Working Paper, Hong Kong University of Science and Technology.
- Francis, J., Wang, D., 2008. The joint effect of investor protection and big 4 audits on earnings quality around the world. *Contemporary Accounting Research* 25 (Spring), 157–191.
- Graham, J.R., Harvey, C.R., Rajgopal, S., 2005. The economic implications of corporate financial reporting. *Journal of Accounting and Economics* 40 (1–3), 3–73.
- Greene, W.H., 2003. *Econometric Analysis*. Pearson Education, New Jersey.
- Heckman, J.J., 1979. Sample selection bias as a specification error. *Econometrica* 47, 153–162.
- Hindo, B., 2003. Audit clients get the heave-ho. *Business Week* (December), 7.
- Huang, R.D., Li, H., 2009. Does the market dole out collective punishment? An empirical analysis of industry, geography, and Arthur Andersen's reputation. *Journal of Banking and Finance* 33 (7), 1255–1265.
- Johnson, P., Jamal, K., Berryman, R., 1991. Effects of framing on auditor decisions. *Organization Behavior and Human Decision Processes* 50, 75–105.
- Kanagaretnam, K., Lobo, G., Yang, D., 2004. Joint tests of signaling and income smoothing through bank loan loss provisions. *Contemporary Accounting Research* (Winter), 843–884.
- Kanagaretnam, K., Krishnan, G., Lobo, G.J., 2009. Is the market valuation of banks' loan loss provision conditional on auditor reputation? *Journal of Banking and Finance* 33 (6), 1039–1047.
- Krishnan, G., 2003. Audit quality and the pricing of discretionary accruals. *Auditing: A Journal of Practice & Theory* (March), 109–126.
- Krishnan, G., 2005. The association between Big 6 auditor industry expertise and the asymmetric timeliness of earnings. *Journal of Accounting, Auditing & Finance* (Summer), 209–228.
- Kwok, C., Tadesse, S., 2006. National culture and financial systems. *Journal of International Business Studies* 39, 1178–1196.
- Kwon, S.Y., Lim, C.Y., Tan, P.M.S., 2007. Legal systems and earnings quality: The role of auditor industry specialization. *Auditing: A Journal of Practice & Theory* 26, 25–55.
- LaPorta, R., Lopez-de-Silanes, F., Shleifer, A., Vishny, R.W., 1998. Law and finance. *The Journal of Political Economy* 106 (6), 1113–1155.
- Leuz, C., Nanda, D., Wysocki, P., 2003. Earnings management and investor protection: An international comparison. *Journal of Financial Economics* 69, 505–527.
- Maddala, G.S., 1977. *Econometrics*. McGraw-Hill, New York.
- Neal, T., Riley Jr., R., 2004. Auditor industry specialist research design. *Auditing: A Journal of Practice & Theory* 23, 169–177.
- O'Keefe, T., King, R., Gaver, K., 1994. Audit fees, industry specialization, and compliance with GAAS reporting standards. *Auditing: A Journal of Practice & Theory* (Fall), 41–55.
- Petersen, M., 2009. Estimating standard errors in finance panel data sets: Comparing approaches. *Review of Financial Studies* 22 (1), 435–480.
- Purda, L., 2008. Risk perception and the financial system. *Journal of International Business Studies* 37, 227–247.
- Rogers, W., 1993. Regression standard errors in clustered samples. *STATA Technical Bulletin* 13, 19–23.
- Shen, C., Chih, H., 2005. Investor protection, prospect theory, and earnings management: An international comparison of the banking industry. *Journal of Banking and Finance* 29 (10), 2675–2697.
- Teoh, S., Wong, T., 1993. Perceived auditor quality and earnings response coefficient. *The Accounting Review* 68 (April), 346–367.
- Wahlen, J.M., 1994. The nature of information in commercial bank loan loss disclosures. *The Accounting Review* (July), 455–478.
- Williams, R., 2000. A note on robust variance estimation for clustered-correlated data. *Biometrics* 56, 645–646.
- Wright, S., Wright, A., 1997. The effect of industry experience on hypothesis generation and audit planning decisions. *Behavioral Research in Accounting* 9, 273–294.